

The distribution of the uncitedness factor and its functional relation with the impact factor

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The uncitedness factor (U) of a journal is its fraction of uncited articles. Given a set of journals (e.g. in a field) we can determine the rank-order distribution of these uncitedness factors. Hereby we use the Central Limit Theorem which is valid for uncitedness factors since they are fractions, hence averages.

A similar result was proved earlier for the impact factors (IF) of a set of journals. Here we combine the two rank-order distributions, hereby eliminating the rank, yielding the functional relation between the impact factor (IF) and the uncitedness factor (U).

The rank-order
distribution of
the impact factor IF

In an earlier paper we studied the rank-order distribution of impact factors (IF) say of a set of journals (e.g. in a field). Remarking that IFs are averages (average number of citations per article in a journal) we can use the Central Limit Theorem (CLT) for the distribution of IFs over these journals:

The normal (or Gaussian) distribution

$$\varphi(\mathbf{x}) = A e^{-\frac{(\mathbf{x}-\mu)^2}{2\sigma^2}}$$

where $\mathbf{x} = \mathbf{IF} \geq 0$ and where the constant A is such that

$$\int_0^{\infty} \varphi(\mathbf{x}) d\mathbf{x} = T$$

the total number of journals.

Ranking the journals in decreasing order of their IFs we have that

$$r = \int_x^{\infty} \varphi(y) dy$$

with $x=IF(r)$, the IF of the journal at rank r (continuous argument). Denoting

$$F(x) = \int_0^x \varphi(y) dy$$

, the cumulative normal distribution, we have

$$\mathbf{r} = \mathbf{T} - \int_0^{\mathbf{x}} \varphi(\mathbf{y}) d\mathbf{y}$$

$$\mathbf{r} = \mathbf{T} - \mathbf{F}(\mathbf{x})$$

Hence, denoting by \mathbf{F}^{-1} the inverse of the injective function \mathbf{F} ,

$$\mathbf{IF}(\mathbf{r}) = \mathbf{x} = \mathbf{F}^{-1}(\mathbf{T} - \mathbf{r})$$

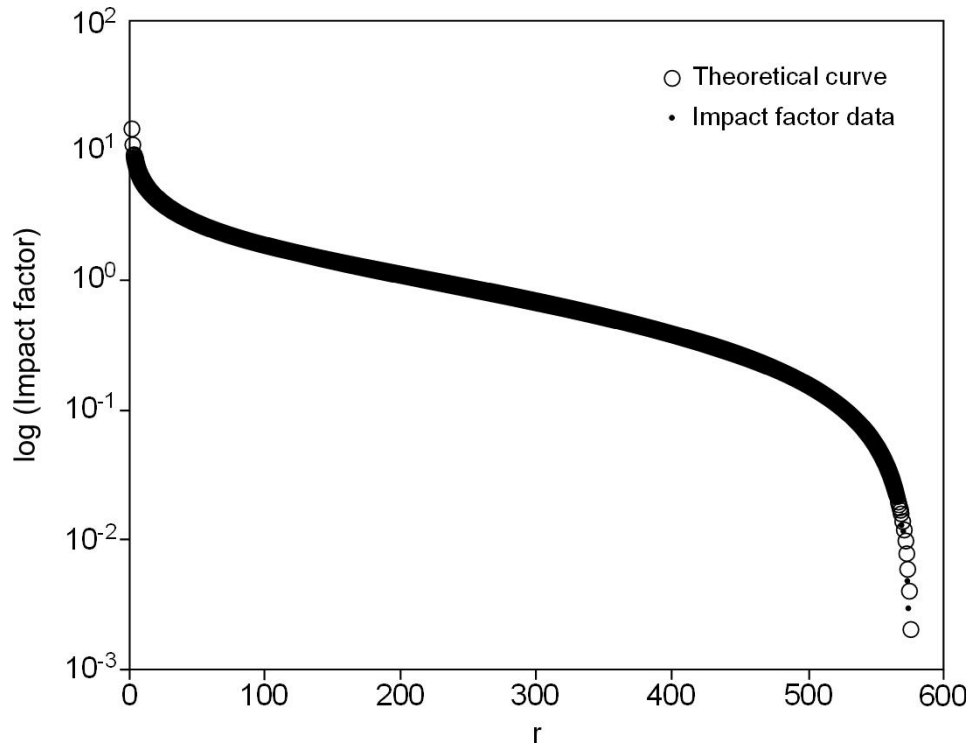
yielding the desired distribution.

Theorem: The function

$$f(r) = \ln(\text{IF}(r)) = \ln(F^{-1}(T - r))$$

is strictly decreasing, first, on an interval $[0, r_0]$ convexly and on the interval $[r_0, T]$ concavely.

Furthermore we have that $\text{IF}(r_0) > \mu$ in the inflection point $(r_0, \ln(\text{IF}(r_0)))$. This is proved based on properties of the Gaussian curve.



Rank-order distribution of $\ln(\text{IF})$. Reprinted from Mansilla, Köppen, Cocho and Miramontes (2007) with kind permission from Elsevier.

Note that the inflection point is indeed in a rank such that $\text{IF}(r_0) > \mu$ (= average of the IFs, i.e. of the Gaussian distribution of IF), i.e. the convex part is “smaller” than the concave part.

We can also show:

Theorem: The function

$$g(\mathbf{r}) = \text{IF}(\mathbf{r}) = F^{-1}(T - \mathbf{r})$$

is strictly decreasing, first on an interval $[0, r_1]$ convexly and on the interval $[r_1, T]$ concavely. Furthermore we have that $\text{IF}(r_1) = \mu$ in the inflection point $(r_1, \text{IF}(r_1))$.

The rank-order
distribution of the
uncitedness factor U

Similar as what we did for the IF, we suppose that U is distributed according to a normal (Gaussian) distribution:

$$\psi(y) = B e^{-\frac{(y-\mu')^2}{2\sigma'^2}}$$

where B is such that

$$\int_0^1 \psi(y) dy = T$$

(note that $y \in [0,1]$ being an uncitedness factor).

Indeed, also for U we can apply the CLT since U is a fraction: the fraction of the papers in the journal which are uncited.

Now we rank the journals in increasing order r of their uncitedness factors U . The defining relation for $U(r)$ is now, evidently,

$$r = \int_0^y \psi(y') dy'$$

where $y = U(r)$. Denote $G(y)$ this integral (which is the cumulative normal distribution) we simply obtain

$$r = G(U)$$

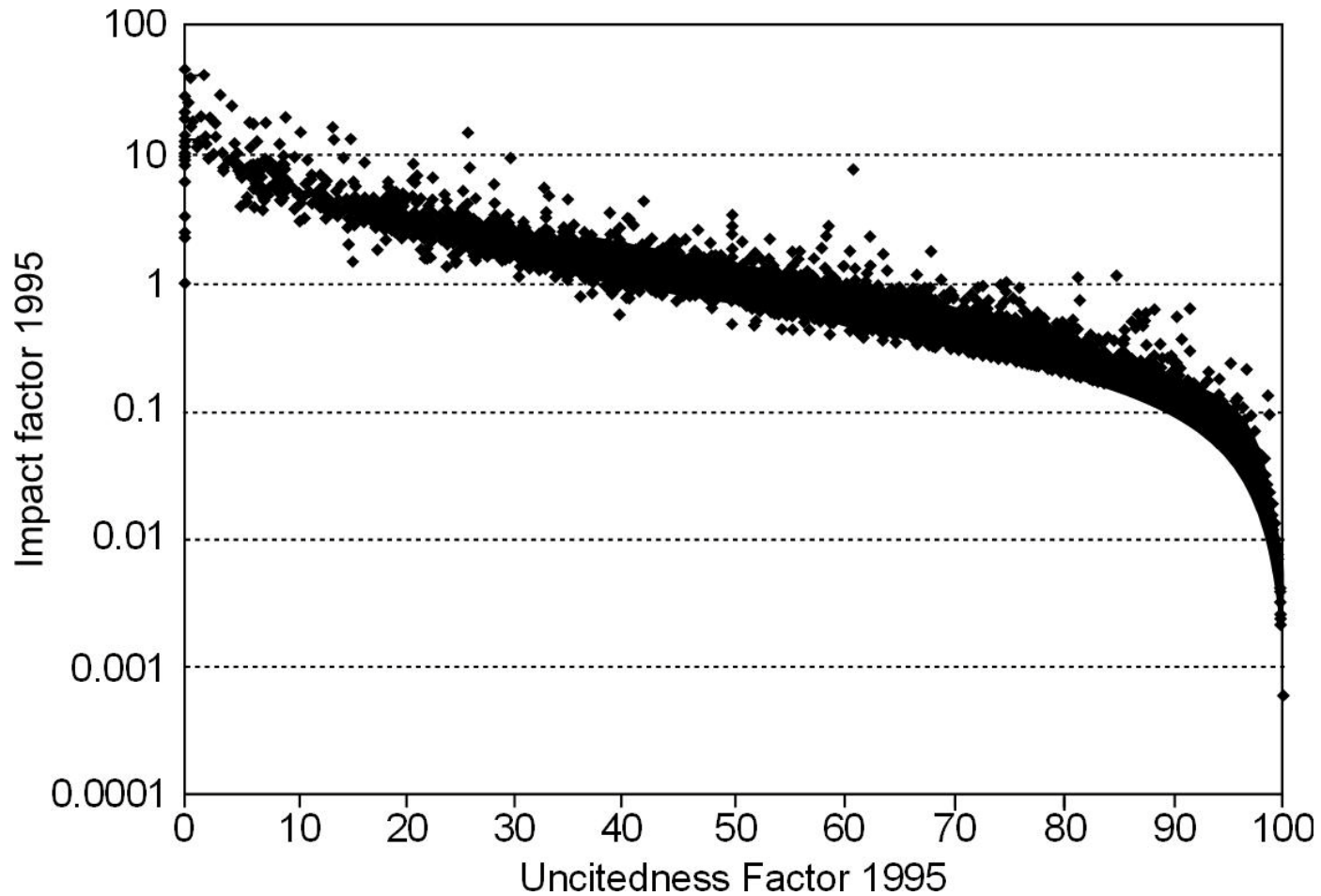
hence

$$U(r) = G^{-1}(r)$$

where G^{-1} is the inverse of the injective function G .

Using properties of the Gaussian distribution we can prove that the function $U(r)$ is first concavely increasing followed by a convex increase. The inflection point is in $(r, U(r))$ where $U(r) = \mu'$ the average of the U s, i.e. of the Gaussian distribution of U .

The functional relation
between the impact
factor IF and the
uncitedness factor U.



$\ln(\text{IF}(U))$, experimentally: SCI (all fields), 1995, van Leeuwen and Moed (2005)

In this section we make one assumption: the rank r occurring in the IF-distribution is the same as the rank r in the U-distribution. This is model-theoretically acceptable since this assumption is equivalent by supposing that IF decreases with U (note that $IF(r)$ decreases in r and $U(r)$ increases in r).

We proved already

$$r = T - F(x)$$

with $x=IF(r)$ and

$$r = G(y)$$

with $y = U(r)$.

$$\mathbf{T} - \mathbf{F}(\mathbf{x}) = \mathbf{G}(\mathbf{y})$$

$$\mathbf{F}(\mathbf{x}) = \mathbf{T} - \mathbf{G}(\mathbf{y})$$

$$\mathbf{x} = \mathbf{F}^{-1}(\mathbf{T} - \mathbf{G}(\mathbf{y}))$$

$$\mathbf{IF}(\mathbf{U}) = \mathbf{IF} = \mathbf{F}^{-1}(\mathbf{T} - \mathbf{G}(\mathbf{U}))$$

involving two different Gaussian distributions.

$$\mathbf{IF}(\mathbf{U}) = \mathbf{F}^{-1}(\mathbf{T} - \mathbf{G}(\mathbf{U}))$$

$$\mathbf{IF}'(\mathbf{U}) = \frac{-\mathbf{G}'(\mathbf{U})}{\mathbf{F}'(\mathbf{F}^{-1}(\mathbf{T} - \mathbf{G}(\mathbf{U})))}$$

$$= \frac{-\psi(\mathbf{U})}{\varphi(\mathbf{IF}(\mathbf{U}))} < \mathbf{0}$$

$$\mathbf{U} = \mathbf{0} \Rightarrow \mathbf{IF}(\mathbf{0}) = \mathbf{F}^{-1}(\mathbf{T}) = +\infty$$

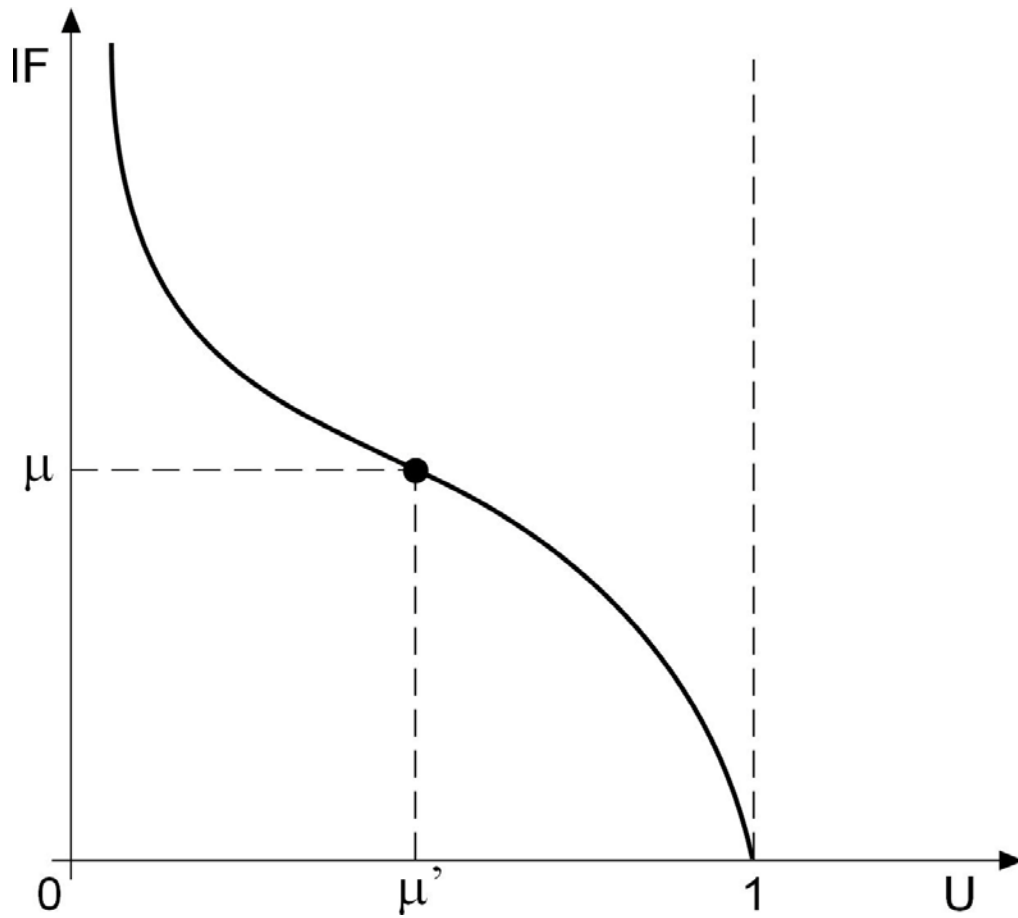
$$\mathbf{U} = \mathbf{1} \Rightarrow \mathbf{IF}(\mathbf{1}) = \mathbf{F}^{-1}(\mathbf{0}) = \mathbf{0}$$

$$\mathbf{IF}''(\mathbf{U}) = -\frac{1}{\left(\mathbf{F}'(\mathbf{F}^{-1}(\mathbf{T} - \mathbf{G}(\mathbf{U})))\right)^2} \frac{\mathbf{F}''(\mathbf{F}^{-1}(\mathbf{T} - \mathbf{G}(\mathbf{U})))}{\mathbf{F}'(\mathbf{F}^{-1}(\mathbf{T} - \mathbf{G}(\mathbf{U})))} (-\mathbf{G}'(\mathbf{U}))(-\mathbf{G}'(\mathbf{U}))$$

$$-\frac{1}{\mathbf{F}'(\mathbf{F}^{-1}(\mathbf{T} - \mathbf{G}(\mathbf{U})))} \mathbf{G}''(\mathbf{U})$$

$$\mathbf{IF}''(\mathbf{U}) = -\frac{1}{\varphi(\mathbf{IF})^3} \varphi(\mathbf{IF}) \left(-\frac{\mathbf{IF} - \mu}{\sigma^2}\right) \psi^2(\mathbf{U}) - \frac{1}{\varphi(\mathbf{IF})} \psi(\mathbf{U}) \left(-\frac{\mathbf{U} - \mu'}{\sigma'^2}\right)$$

$\Rightarrow \mathbf{IF}''(\mu') = 0$ in the point (μ, μ') on the curve
 $(\mathbf{IF} = \mu, \mathbf{U} = \mu')$



$$f(U) = \ln(\text{IF}(U)) = \ln(F^{-1}(T - G(U)))$$

Now $f''(U)$ has the sign of

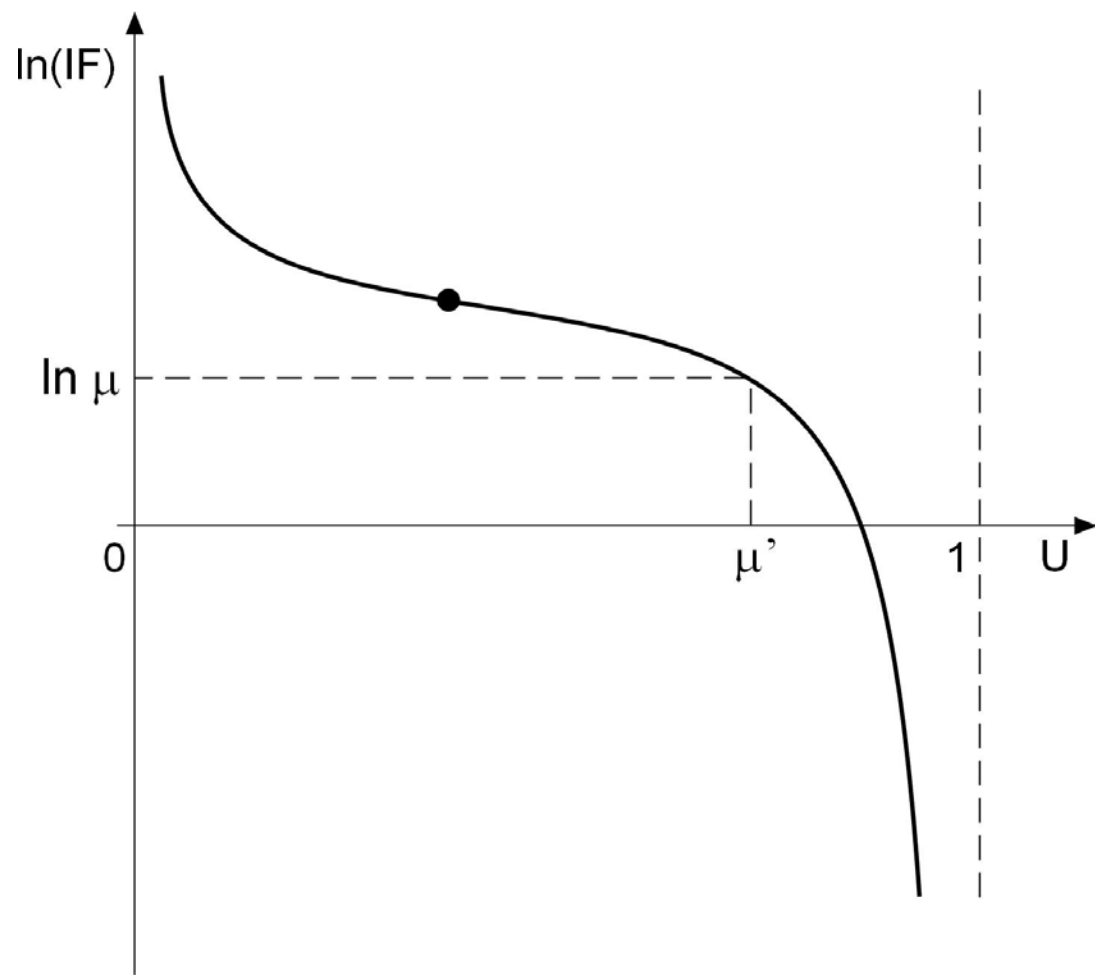
$$-\frac{\psi(U)}{\text{IF}(U)\varphi(\text{IF}(U))} + \frac{\psi(U)\left(\frac{\text{IF}(U) - \mu}{\sigma^2}\right)}{\varphi(\text{IF}(U))} + \frac{U - \mu'}{\sigma'^2}$$

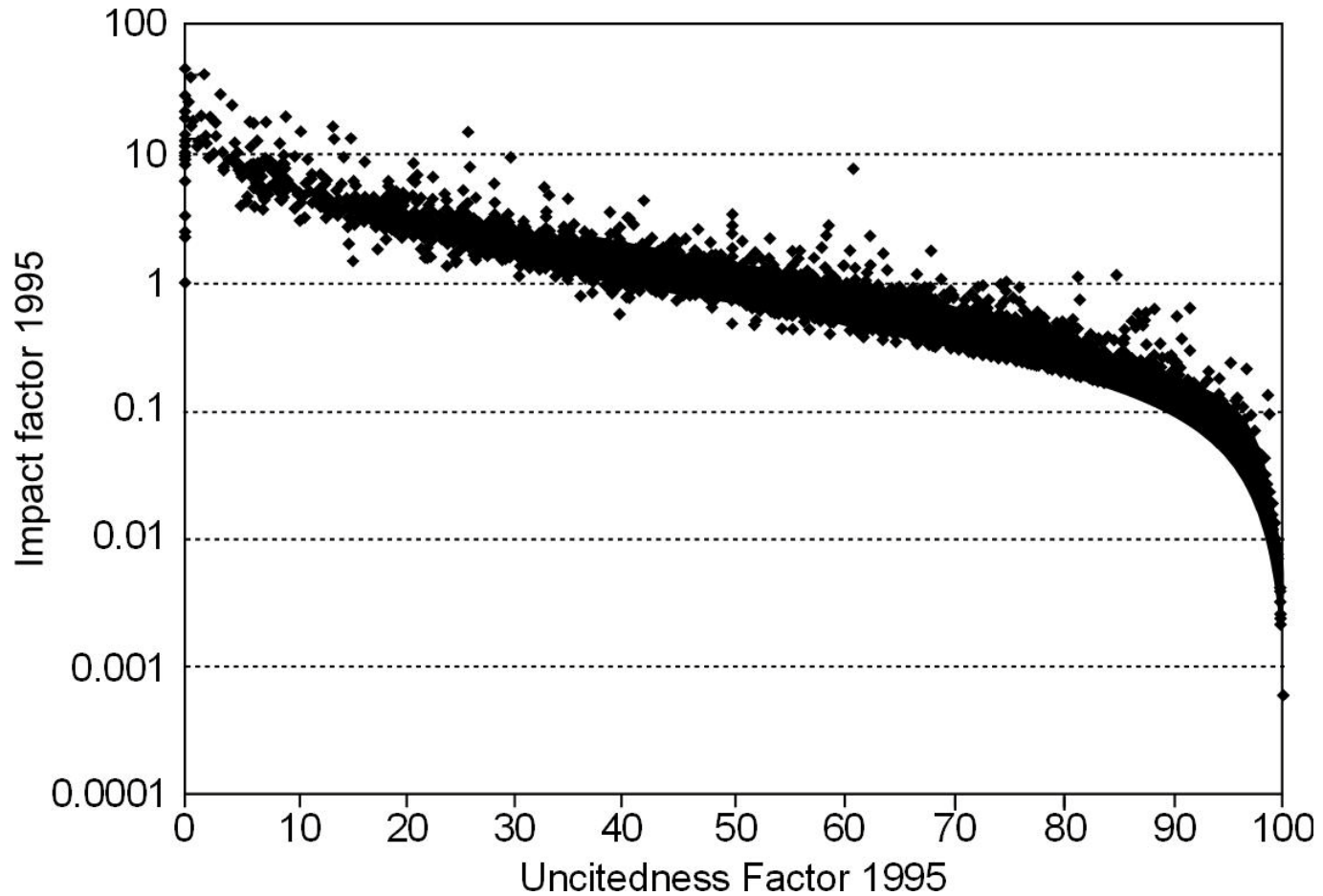
We now see that in the point $(\mu', \ln \mu)$ (i.e. for $U = \mu'$ and $\text{IF} = \mu$) that $f''(U)$ has the sign of

$$\frac{-\psi(U)}{\text{IF}(U)\varphi(\text{IF}(U))} < 0$$

hence strictly negative.

This implies, when compared to $\text{IF}(U)$ (where in this point one had an inflection point), that $\ln(\text{IF}(U))$ already concavely decreasing in this point.





$\ln(\text{IF}(U))$, experimentally: SCI (all fields), 1995, van Leeuwen and Moed (2005)